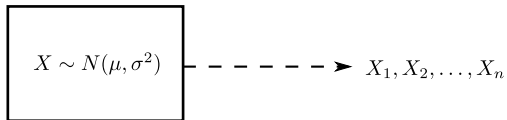


## Lecture 30: Confidence Intervals for $\sigma^2$

Today we will discuss the material in Section 7.4.

Let  $X_1, X_2, \dots, X_n$  be a random sample from a normal population with mean  $\mu$  and variance  $\sigma^2$ .



In this lecture we want to construct a  $100(1 - \alpha)\%$  confidence for  $\sigma^2$ . We recall that  $S^2$  is a point estimator for  $\sigma^2$ .

What is new here is that we are going to note a “*multiplicative confidence interval*”.

Here is the idea. *We want a random interval that has the point estimator  $S^2$  in the interior*

Now given a number  $x$  there are two ways to make an interval  $I(x)$  that has  $x$  in its interior.

### 1. The additive method

Choose two positive numbers  $c_1$  and  $c_2$ . Put  $I(x) = (x - c_1, x + c_2)$ .

### 2. The multiplicative method

Choose a number  $c_1 < 1$  and another number  $c_1 > 1$ . Put

$$I(x) = (c_1 x, c_2 x).$$

We will use the second method now. The clue to why we do this is that  $S^2 > 0$ . First we need to know the probability distribution of the point estimator  $S^2$ . We have already seen this

#### Theorem A (pg 278)

$$V = \left( \frac{n-1}{\sigma^2} \right) S^2 \sim \chi^2(n-1) \quad (*)$$

*Now we can give the confidence interval.*

#### Theorem B

*The random interval  $\left( \frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2, \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2 \right)$  is a  $100(1-\alpha)\%$  confidence random interval for the population variance  $\sigma^2$  from a normal population.*

## Remark

*It must be true (see page 2) that*

$$c_1 = \frac{n-1}{\chi_{\alpha/2, n-1}^2} < 1 \quad \text{and}$$
$$c_2 = \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} > 1.$$

*I have never checked this.*

*Now we prove Theorem B. We must prove*

$$P\left(\sigma^2 \in \left(\frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2, \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2\right)\right) = 1$$
$$\text{LHS} = P\left(\frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2 < \sigma^2, \sigma^2 < \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2\right)$$

## Remark

Now we manipulate the two resulting inequalities to get  $V$  so we can sue (\*)

$$P\left(\frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2 < (\sigma^2), (\sigma^2) < \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2\right)$$

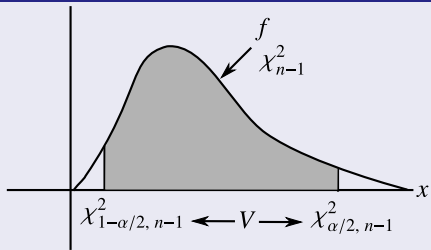
Swap and make a  $V$

$$\begin{aligned} &= P\left(\frac{n-1}{\sigma^2} S^2 < \chi_{\alpha/2, n-1}^2, \chi_{1-\alpha/2, n-1}^2 < \frac{n-1}{\sigma^2} S^2\right) \\ &= P\left(V < \chi_{\alpha/2, n-1}^2, \chi_{1-\alpha/2, n-1}^2 < V\right) \\ &= P\left(\chi_{1-\alpha/2, n-1}^2 < V < \chi_{\alpha/2, n-1}^2\right) \end{aligned}$$

MAKE A PICTURE

= the shaded area

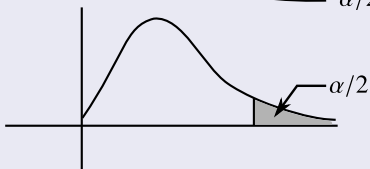
## Remark (Cont.)



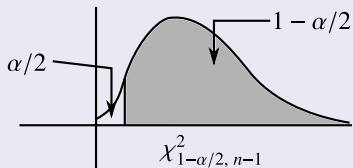
$$= 1 - (\triangleleft + \triangleright)$$

$\alpha/2$

Now



and



$$= 1 - (\alpha/2 + \alpha/2) = 1 - \alpha$$

## Question

Why do we need the strange  $\chi^2_{1-\alpha/2}$ ,  $n - 1$ ? This is because the  $\chi^2$  density curve does not have the symmetry that the  $z$ -density and  $t$ -densities did. In all three cases we need something that cut off  $\alpha/2$  on the *left* under the density curve so  $1 - \alpha/2$  on the *right*. For the  $z$ -curve  $-z_{\alpha/2}$  did the job.

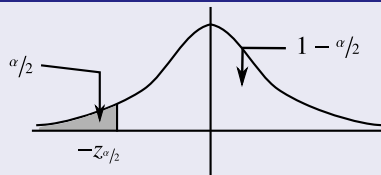
In other words



## Lemma

$$Z_{1-\alpha/2} = -Z_{\alpha/2}$$

## Proof.



so  $-Z_{\alpha/2}$  cuts off  $1 - \alpha/2$  to the right to  $-Z_{\alpha/2} = Z_{1-\alpha/2}$

□

## The Upper-Tailed $100(1 - \alpha)\%$ Confidence Interval for $\sigma^2$

### Theorem

$\left( \frac{n-1}{\chi_{\alpha, n-1}^2} S^2, \infty \right)$  is a  $100(1 - \alpha)\%$  confidence interval for  $\sigma^2$

### Proof.

If could be on the final - do it yourself. □

### Remark

*As used we took the lower limit from the two-sided interval and changed  $\alpha/2$  to  $\alpha$ .*

## The Lower-Tailed $100(1 - \alpha)\%$ Confidence Interval for $\sigma^2$

Since  $S^2$  is always positive  $PCS^2 \in (-\infty, 0] = \emptyset$  so the negative axis will not appear.

*Lower tailed multiplication intervals go down to 0 not  $-\infty$ . Another (philosophical) way to look at it is.*

$$\underbrace{\begin{array}{c} \text{additive group of } \mathbb{R} \\ (-\infty, \infty) \end{array}}_{\text{additive world}} \xrightarrow{e^x} \underbrace{\begin{array}{c} \text{multiplicative group of positive} \\ \text{numbers, } (0, \infty) \end{array}}_{\text{multiplicative world}}$$

We are in the multiplicative world.

## Theorem

$\left(0, \frac{n-1}{\chi_{1-\alpha, n-1}^2} S^2\right)$  is a  $100(1-\alpha)\%$  confidence interval for  $\sigma^2$ .

## Proof.

Do it yourself



## Remark

$\left(-\infty, \frac{n-1}{\chi_{1-\alpha, n-1}^2} S^2\right)$  is also a  $100(1-\alpha)\%$  confidence interval for  $\sigma^2$  but the  $(-\infty, 0)$  is “wasted space”, Remember, small intervals or better.

## Confidence Intervals for the standard Deviation

Note that if  $a > 0$ ,  $b > 0$  and  $x > 0$  then

$$a \leq x \leq b \leftrightarrow \sqrt{a} \leq \sqrt{x} \leq \sqrt{b}$$

so

$$\begin{aligned} \frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2 \leq \sigma^2 \leq \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2 \\ \Leftrightarrow \sqrt{\frac{n-1}{\chi_{\alpha/2, n-1}^2}} S \leq \sigma \leq \sqrt{\frac{n-1}{\chi_{1-\alpha/2, n-1}^2}} S \end{aligned}$$

Hence

$$\begin{aligned} \left( \sqrt{\frac{n-1}{\chi_{\alpha/2, n-1}^2}} S < \sigma < \sqrt{\frac{n-1}{\chi_{1-\alpha/2, n-1}^2}} S \right) &= P \left( \frac{n-1}{\chi_{\alpha/2, n-1}^2} S^2 \leq \sigma^2 \leq \frac{n-1}{\chi_{1-\alpha/2, n-1}^2} S^2 \right) \\ &\text{from pg3} \\ &= 1 - \alpha \end{aligned}$$

In other words

$$P\left(\sigma \in \left(\sqrt{\frac{n-1}{\chi_{\alpha/2, n-1}^2}}S, \sqrt{\frac{n-1}{\chi_{1-\alpha/2, n-1}^2}}S\right)\right) = 1 - \alpha$$

and we have

### Theorem

*The random interval*

$$\left(\sqrt{\frac{n-1}{\chi_{\alpha/2, n-1}^2}}S, \sqrt{\frac{n-1}{\chi_{1-\alpha/2, n-1}^2}}S\right)$$

*is a  $100(1 - \alpha)\%$  confidence interval for the standard deviation  $\sigma$  in a normal population.*

## Problem

*Write down the upper and lower-tailed confidence intervals for  $\sigma$ .  
(hint: just take the square notes of the end points of those for  $\sigma^2$ )*