

The Cramér-Rao Lower Bound in the Phase Retrieval Problem

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Notations and Assumptions

Phase Retrievability and Identifiability

- Hilbert space $H = \mathbb{C}^n$, $\hat{H} = H/T^1$, frame $\mathcal{F} = \{f_1, \dots, f_m\} \subset \mathbb{C}^n$ and $\alpha : \hat{H} \rightarrow \mathbb{R}^m$, $\alpha(x) = (|\langle x, f_k \rangle|)_{1 \leq k \leq m}$.
- We assume the frame is *phase retrievable*, i.e., α is injective. Hence $(|\langle x, f_k \rangle|)_{1 \leq k \leq m}$ determine uniquely x up to a global phase factor.

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- We assume the frame is *phase retrievable*, i.e., α is injective. Hence $(|\langle x, f_k \rangle|)_{1 \leq k \leq m}$ determine uniquely x up to a global phase factor.
- Measurement process: $y = (y_k)_{1 \leq k \leq m}$. We assume the distribution of y , $p(y; x)$ depends on $\alpha(x)$ only. For instance:

$$y_k = |\langle x, f_k \rangle + \mu_k|^a + \nu_k, \quad \mu_k \sim \mathbb{CN}(0, \rho^2), \quad \nu_k \sim \mathbb{N}(0, \sigma^2)$$

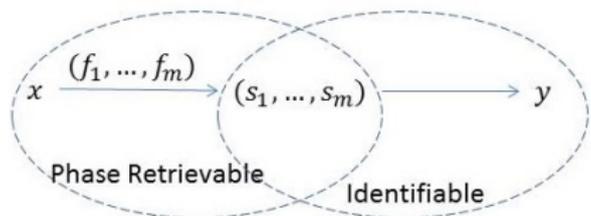
Specifically: $p(y; x) = F(s_1, \dots, s_m, y)$, where $s_k = |\langle x, f_k \rangle|$.

- We assume *identifiability and regularity*: (1) If $\forall y \in \mathbb{R}^m$, $F(s^{[1]}, y) = F(s^{[2]}, y)$ then $s^{[1]} = s^{[2]}$; and, (2) The Fisher Infomatrix $\mathbb{E}\left[\frac{\partial \log(F)}{\partial s_k} \frac{\partial \log(F)}{\partial s_j}\right]$ is continuous and has constant rank on an open neighborhood of the operating point [Rthbrg71].

Problem Statement

FIM vs. CRLB

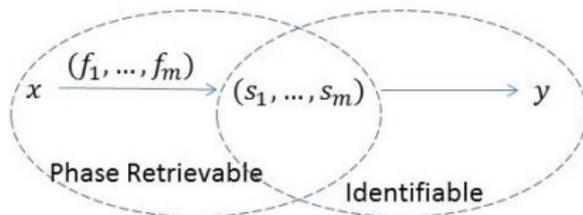
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In previous works we derived various Fisher Information Matrix expressions. We have also derived a Cramér-Rao Lower Bound (CRLB) for a specific estimation model. In this paper we analyze a second identification problem and compare the two CRLBs:

Problem

*The problem is not how to compute the Fisher Information Matrix (FIM).
The problem is how to use FIM, to derive Cramér-Rao Lower Bounds.*

Fisher Info Matrix for the AWGN Model

- For the AWGN model:

$$y_k = |\langle x, f_k \rangle|^2 + \nu_k \quad , \quad 1 \leq k \leq m$$

with $\nu_k \sim \mathcal{CN}(0, \sigma^2)$ i.i.d. the Fisher Information Matrix:

$$\mathbb{I} = \mathbb{E} [(\nabla_x \log p(y; x))(\nabla_x \log p(y; x))^*]$$

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$$\mathbb{I} = \mathbb{E} [(\nabla_x \log p(y; x))(\nabla_x \log p(y; x))^*]$$

- $\mathbb{I}^{AWGN, real}(x) = \frac{4}{\sigma^2} \sum_{k=1}^m |\langle x, f_k \rangle|^2 f_k f_k^T = \frac{4}{\sigma^2} \sum_{k=1}^m (f_k f_k^T) x x^T (f_k f_k^T)$

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- $\mathbb{I}^{AWGN, cplx}(x) = \frac{4}{\sigma^2} \sum_{k=1}^m \Phi_k \xi \xi^* \Phi_k$ [Bal13, BCMN13] with $\Phi_k \in \mathbb{R}^{2n \times 2n}$ and $\xi \in \mathbb{R}^{2n}$.

FIM for Non-AWGN

- Consider the Non-AWGN model:

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- The likelihood function:

$$p(y; x) = \frac{1}{\rho^{2m}} \exp \left\{ -\frac{1}{\rho^2} \left(\sum_{k=1}^m y_k + \sum_{k=1}^m |\langle x, f_k \rangle|^2 \right) \right\} \prod_{k=1}^m I_0 \left(\frac{2|\langle x, f_k \rangle| \sqrt{y_k}}{\rho^2} \right)$$

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- Realification: $x \mapsto \xi = [\text{real}(x) \ \text{imag}(x)]^T$ and $|\langle x, f_k \rangle| = \sqrt{\langle \Phi_k \xi, \xi \rangle}$ where Φ_k is a rank-2 replacing $f_k f_k^*$.

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FIM for Non-AWGN

Theorem (Bal15)

The Fisher Information Matrix for the Non-AWGN model is given by

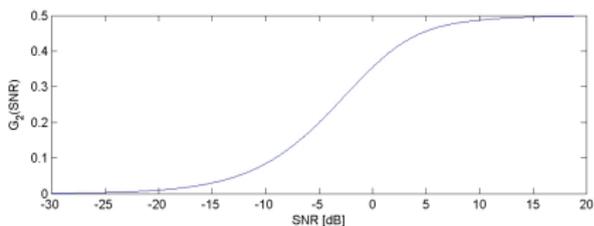
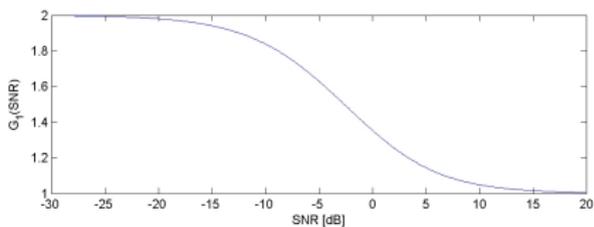
$$\begin{aligned} \mathbb{I}(\xi) &= \frac{4}{\rho^4} \sum_{k=1}^m \left(G_1 \left(\frac{\langle \Phi_k \xi, \xi \rangle}{\rho^2} \right) - 1 \right) \Phi_k \xi \xi^* \Phi_k \\ &= \frac{4}{\rho^2} \sum_{k=1}^m G_2 \left(\frac{\langle \Phi_k \xi, \xi \rangle}{\rho^2} \right) \frac{1}{\langle \Phi_k \xi, \xi \rangle} \Phi_k \xi \xi^* \Phi_k \end{aligned}$$

where

$$G_1(a) = \frac{e^{-a}}{8a^3} \int_0^\infty \frac{I_1^2(t)}{I_0(t)} t^3 e^{-\frac{t^2}{4a}} dt, \quad G_2(a) = a(G_1(a) - 1)$$

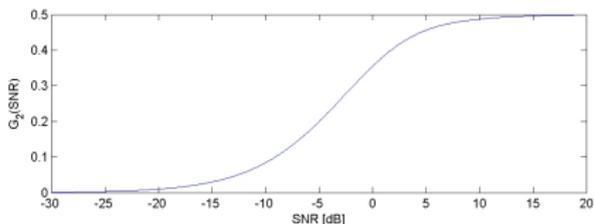
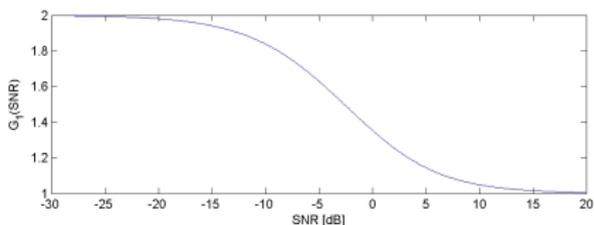
FIM for Non-AWGN

Asymptotic Regimes



FIM for Non-AWGN

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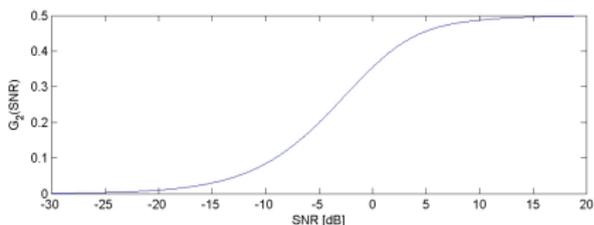
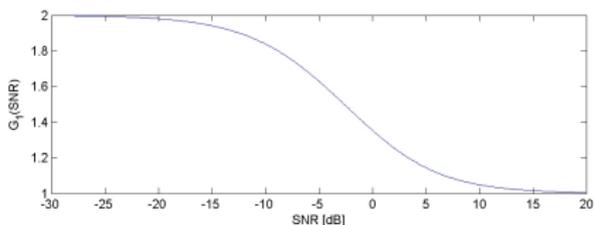


Form 1: Low SNR

$$\begin{aligned} \mathbb{I}(\xi) &= \\ & \frac{4}{\rho^4} \sum_{k=1}^m \left(G_1 \left(\frac{\langle \Phi_k \xi, \xi \rangle}{\rho^2} \right) - 1 \right) \Phi_k \xi \xi^* \Phi_k \\ & \approx \frac{4}{\rho^4} \sum_{k=1}^m \Phi_k \xi \xi^* \Phi_k \end{aligned}$$

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Form 1: Low SNR

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Form 2: High SNR

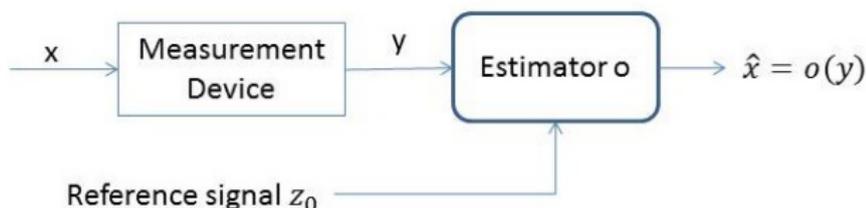
$$\begin{aligned} \mathbb{I}(\xi) &= \\ & \frac{4}{\rho^2} \sum_{k=1}^m G_2 \left(\frac{\langle \Phi_k \xi, \xi \rangle}{\rho^2} \right) \frac{1}{\langle \Phi_k \xi, \xi \rangle} \Phi_k \xi \xi^* \Phi_k \\ & \approx \frac{2}{\rho^2} \sum_{k=1}^m \frac{1}{\langle \Phi_k \xi, \xi \rangle} \Phi_k \xi \xi^* \Phi_k \end{aligned}$$

Setup 1: Reference signal based estimation

In the first setup we fix a reference unit-norm signal $z_0 \in \mathbb{C}^n$. The unknown (to-be-estimated) signal x is assumed to come from set:

$$V_{z_0} = \{x \in \mathbb{C}^n : \text{imag}(\langle x, z_0 \rangle) = 0, \text{real}(\langle x, z_0 \rangle) > 0\}.$$

The estimator has access to the reference signal z_0 :

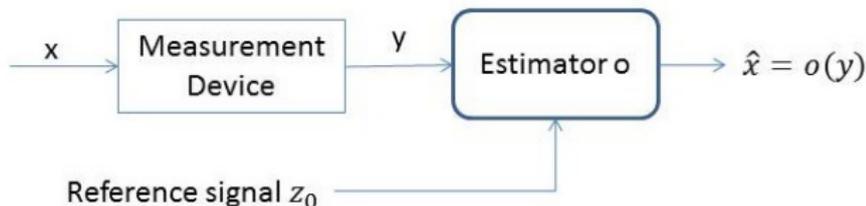


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Let $\mathcal{V}_{\zeta_0} = \{\xi \in \mathbb{R}^{2n}, \langle \xi, \zeta_0 \rangle \geq 0, \langle \xi, J\zeta_0 \rangle = 0\}$, $\mathcal{E}_{\zeta_0} = \text{span}_{\mathbb{R}}(\mathbb{V}_{\zeta_0})$ with $\zeta_0 = [\text{real}(z_0) \ \text{imag}(z_0)]^T$. The estimator $o : \mathbb{R}^m \rightarrow \mathcal{E}_{\zeta_0}$ is *unbiased* if $\mathbb{E}[o(y); \xi] = \xi$ for every $x \in V_{z_0}$, with $\xi = [\text{real}(x); \text{imag}(x)]$.

Setup 1: Positive correlation with a reference signal

The CRL Bound

Let $\Pi_\eta = 1 - \frac{1}{\|\eta\|^2} J\eta\eta^T J^T$ and $L = I - \frac{1}{\langle \xi, \zeta_0 \rangle} J\zeta_0\xi^T J^T$, with J the symplectic form matrix $[0, -I; I, 0]$.

Theorem

Assume the measurement model $y = (y_k)_{1 \leq k \leq m}$ where the likelihood function $p(y; x) = F(|\langle x, f_1 \rangle|, \dots, |\langle x, f_m \rangle|, y)$ is identifiable and regular. Then the covariance of any unbiased estimator $\omega : \mathbb{R}^m \rightarrow \mathcal{E}_{\zeta_0}$ is bounded below by

$$\text{Cov}[\omega(y); \xi] \geq (\Pi_{z_0} \mathbb{I}(\xi) \Pi_{z_0})^\dagger = L^T (\mathbb{I}(\xi))^\dagger L.$$

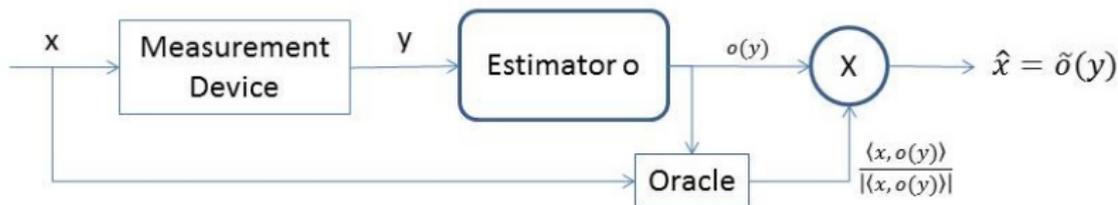
In particular: $\mathbb{E}[\|\omega(y) - \xi\|^2; \xi] \geq \text{trace} \left\{ (\Pi_{z_0} \mathbb{I}(\xi) \Pi_{z_0})^\dagger \right\} = \text{trace}(\mathbb{I}(\xi))^\dagger + \frac{\|\xi\|^2}{|\langle \xi, \zeta_0 \rangle|^2} \langle (\mathbb{I}(\xi))^\dagger J\zeta_0, J\zeta_0 \rangle.$

Remark: First inequality was derived in 2015 paper; the second equality is new.

Setup 2: Oracle-based signal estimation

Consider now a different setup, where $x \in \mathbb{C}^n$ is unconstrained and the estimation is performed in two stages: (i) the first stage returns a "class" estimate through $o : \mathbb{R}^m \rightarrow \mathbb{C}^n$; (ii) in the second stage, an oracle provides the optimal global phase $\frac{\langle x, o(y) \rangle}{|\langle x, o(y) \rangle|}$. Thus, the overall estimator:

$$\tilde{o} : \mathbb{R}^m \rightarrow \mathbb{C}^n, \quad \tilde{o}(y) = o(y) \frac{\langle x, o(y) \rangle}{|\langle x, o(y) \rangle|}.$$



The estimator is *unbiased* if $\mathbb{E}[\tilde{o}(y); x] = x$ for every $x \in \mathbb{C}^n$.

Setup 2: Oracle-based signal estimation

The CRL Bound

Theorem

Assume the measurement model $y = (y_k)_{1 \leq k \leq m}$ where the likelihood function $p(y; x) = F(|\langle x, f_1 \rangle|, \dots, |\langle x, f_m \rangle|, y)$ is identifiable and regular. Let $\tilde{o} : \mathbb{R}^m \rightarrow \mathbb{C}^n$ be an unbiased estimator in Setup 2 (Oracle-based estimator). Denote by $\omega(y) = [\text{real}(o(y)); \text{imag}(o(y))]$ and $\tilde{\omega}(y) = [\text{real}(\tilde{o}(y)); \text{imag}(\tilde{o}(y))]$. Then for any $\xi = [\text{real}(x); \text{imag}(x)] \neq 0$,

$$\text{Cov}[\tilde{\omega}(y); \xi] \geq (I - \Delta)(\mathbb{I}(\xi))^\dagger (I - \Delta)$$

where $\Delta =$

$$\mathbb{E} \left[\frac{(\langle \omega, J\xi \rangle)^2}{((\langle \omega, \xi \rangle)^2 + (\langle \omega, J\xi \rangle)^2)^{3/2}} \omega \omega^T + \frac{\langle \omega, \xi \rangle \langle \omega, J\xi \rangle}{((\langle \omega, \xi \rangle)^2 + (\langle \omega, J\xi \rangle)^2)^{3/2}} (J\omega \omega^T + \omega \omega^T J^T) + \frac{(\langle \omega, \xi \rangle)^2}{((\langle \omega, \xi \rangle)^2 + (\langle \omega, J\xi \rangle)^2)^{3/2}} J\omega \omega^T J^T \right]$$

and satisfies $\Delta = \Delta^T \geq I - \Pi_\xi \geq 0$, $\Delta J\xi = J\xi$ and $\Delta\xi = 0$.

Conclusions and Open Questions

We obtained Cramér-Rao (type) Lower Bounds for two setups:

- 1 Positive Correlation with a reference signal: CRLB has a simple form.
- 2 Oracle-based global phase: CRLB seems very complicated, and estimator dependent. (Remark: Estimator dependency is known for other classes of estimators)

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We obtained Cramér-Rao (type) Lower Bounds for two setups:

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- ② Oracle-based global phase: CRLB seems very complicated, and estimator dependent. (Remark: Estimator dependency is known for other classes of estimators)

Open Question: Which of the two CRL bounds is smaller?

Intuitively, Oracle-based estimator seems to have more information than the reference signal based estimator. But is this true/quantifiable?

Easy case: $CRLB_{Setup\ 1} \rightarrow \infty$ as $\xi \perp \zeta_0$.

Thank you! Merci!

Questions?

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