

Lecture 8: Calibration of SIR Models

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SIR Model Calibration

Recall the model:

$$\begin{cases} \frac{dS}{dt} = -\beta S \frac{I}{N}, & S(0) \\ \frac{dI}{dt} = \beta S \frac{I}{N} - \alpha I, & I(0) \\ \frac{dR}{dt} = \alpha I, & R(0) \end{cases}$$

with the sub-compartments $X(t) = (1 - \gamma)R(t)$ for “recovered” and $Y(t) = \gamma R(t)$ for deaths.

Before making useful predictions (testing), the model has to be calibrated. For calibration and testing we are using two pieces of measured data: the *cumulative detected infections*, $\{V(0), \dots, V(T_{max})\}$, and the time series of *cumulative deaths*, $\{Y(0), \dots, Y(T_{max})\}$. The cumulative detected infections will have to be converted into infection rates $\{I(0), \dots, I(T_{max})\}$. Typically, the cumulative detected infections *undercounts* the true number of infections, hence $I(t) \approx \rho I_{true}(t)$ for some $\rho \leq 1$. Note that, if we know γ and N we can compute $R(0) = \frac{Y(0)}{\gamma}$ and $S(0) = N - I(0) - R(0)$. At the onset of an infectious disease it is likely that $Y(0) = 0$ and $I(0)$ can be neglected in which case, $S(0) = N$ (regardless of γ). This approximation may hold for a certain interval of time, but it certainly fails after significant time has passed since the onset.

The LSE

The least-squares estimator (LSE) finds parameters α, β, γ , and ρ that minimize:

$$\begin{aligned} \text{minimize } & J(\alpha, \beta; \gamma, \rho) = c_I \sum_{t=0}^{T_{max}} (I(t) - \rho I_{sim}(t))^2 + c_Y (Y(t) - \gamma R_{sim}(t))^2 \\ & \alpha, \beta, \gamma, \rho \geq 0 \\ & \gamma, \rho \leq 1 \end{aligned}$$

where $(S_{sim}(t), I_{sim}(t), R_{sim}(t))$ are given by a numerical solver of the SRI model with parameters (α, β, γ) and total population N initialized at $(S(0), I(0), R(0))$. ρ is the undercounting factor, and $c_I, c_Y \geq 0$ are weights.

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First, the initial time $t_0 = 0$ is chosen once a significant number of infections occurred.

The l^p Estimator

The l^p estimator (LPE) finds parameters $\alpha, \beta, \gamma, \rho$ that minimize $I_p(\alpha, \beta, \gamma, \rho)$ defined as follows. For $1 \leq p < \infty$:

$$\begin{aligned} &\text{minimize } I_p(\alpha, \beta; \gamma, \rho) = c_I \sum_{t=0}^{T_{\max}} |I(t) - \rho I_{\text{sim}}(t)|^p + c_Y |Y(t) - \gamma R_{\text{sim}}(t)|^p. \\ &\alpha, \beta, \gamma, \rho \geq 0 \\ &\gamma, \rho \leq 1 \end{aligned}$$

For $p = \infty$,

$$\begin{aligned} &\text{minimize } I_p(\alpha, \beta; \gamma, \rho) = c_I \max_{0 \leq t \leq T_{\max}} |I(t) - \rho I_{\text{sim}}(t)| + c_Y \max_{0 \leq t \leq T_{\max}} |Y(t) - \gamma R_{\text{sim}}(t)|. \\ &\alpha, \beta, \gamma, \rho \geq 0 \\ &\gamma, \rho \leq 1 \end{aligned}$$

As before, $(S_{\text{sim}}(t), I_{\text{sim}}(t), R_{\text{sim}}(t))$ are given by a numerical solver of the SRI model with parameters (α, β, γ) and total population N initialized at $(S(0), I(0), R(0))$. ρ is the undercounting factor, and $c_I, c_Y \geq 0$ are weights.

γ and ρ estimators

We start with the “simpler” problem of estimating γ . Same procedure works for ρ .

Assume $\{V(0), V(1), \dots, V(T_{max})\}$ denotes the cumulative number of detected infections, and $\{Y(0), Y(1), \dots, Y(T_{max})\}$ denote the time series of virus related deaths. It is necessary that

$0 \leq Y(t) \leq V(t) \leq V(T_{max})$ for every $0 \leq t \leq T_{max}$.

Since all infected individuals eventually transit into the “removed” state, $R(t')$, **for calibration purposes** we make the assumption that

$V(t) \approx R(t + \tau)$ for some $\tau > 0$. In fact, τ should be close to $\frac{1}{\alpha}$. In this case we obtain: $Y(t + \tau) \approx \gamma V(t)$. A natural optimization problem is to minimize a norm of the difference between $\gamma V(t)$ and $Y(t + \tau)$. Consider $1 \leq p < \infty$ and define

$$F(\gamma, \tau; p) := \sum_{t=0}^{T_{max}-\tau} |Y(t + \tau) - \gamma V(t)|^p$$

l^p estimators for γ

For $p = \infty$ adjust the definition:

$$F(\gamma, \tau; \infty) := \max_{0 \leq t \leq T_{max} - \tau} |Y(t + \tau) - \gamma V(t)|$$

Consider the optimization problem:

$$\begin{aligned} & \text{minimize} && F(\gamma, \tau; p) \\ & \tau, \gamma \geq 0 \quad , \quad \gamma \leq 1 \end{aligned}$$

for the given calibration data set. In the following we analyze the cases $p = 1, 2, \infty$. In each case, the optimization problem minimizes an l^p norm of the form $\|Y(\cdot + \tau) - \gamma V\|_p$, scaled by the number of terms in each sum.

Good news: The optimization problem is convex. The bad news: Given τ , except for $p = 2$, in the other cases the optimization problem does not have a closed form solution, but can be easily solved.

I^P estimators for γ

The general optimization problem is solved by an iterative algorithm:

Algorithm (Meta-Algorithm for γ estimation)

Inputs: Time series $\{V(0), \dots, V(T_{max})\}$, $\{Y(0), \dots, Y(T_{max})\}$.

Parameters: $p \in [1, \infty]$, τ_{max} .

- ① For each $\tau = 0, 1, 2, \dots, \tau_{max}$ repeat:
 - ① Solve $[Fmin, \gamma_{min}] = \min_{\gamma \in [0, 1]} F(\gamma, \tau; p)$.
 - ② Save vector $F(\tau) = F(\gamma_{min}, \tau; p)$, $optGamma(\tau) = \gamma_{min}$.
- ② Determine the minimum and the minimizer $[minF, \hat{\tau}] = \min(\text{vector}F)$
- ③ Assign $\hat{\gamma} = optGamma(\hat{\tau})$, $F(\hat{\gamma}, \hat{\tau}; p) = minF$.

Outputs: Estimated $\hat{\gamma}, \hat{\tau}$ and minimum value of the objective function $F(\hat{\gamma}, \hat{\tau}; p)$.

Next we analyze the Step 1.1.

The case $p = 2$

The case $p = 2$ is the easiest: it is solved by the least-squares fit with a linear model. Solution of

$$\underset{\gamma}{\text{minimize}} \quad \sum_{t=0}^{T_{\max}-\tau} |Y(t+\tau) - \gamma V(t)|^2$$

is given by:

$$\gamma_c = \frac{\sum_{t=0}^{T_{\max}-\tau} Y(t+\tau)V(t)}{\sum_{t=0}^{T_{\max}-\tau} |V(t)|^2}$$

If the above expression does not belong to $[0, 1]$, the adjust the value to the closest end point:

$$\gamma_{\min} = \begin{cases} 0 & \text{if } \gamma_c < 0 \\ \gamma & \text{if } \gamma_c \in [0, 1] \\ 1 & \text{if } \gamma_c > 1 \end{cases}$$

The case $p = 1$

Solution of optimization problem minimize $\sum_{t=0}^{T_{max}-\tau} |Y(t+\tau) - \gamma V(t)|$:
 γ

Algorithm (The l^1 estimator for γ)

- 1 For each $k = 0, 1, \dots, T_{max} - \tau$ repeat:
 - 1 Compute $r(k) = \frac{Y(k+\tau)}{V(k)}$.
 - 2 If $r(k) \notin [0, 1]$ then discard this value and proceed to the next k .
 - 3 Compute: $f(k) = \sum_{t=0}^{T_{max}-\tau} |Y(t+\tau) - r(k)V(t)|$
- 2 Find the minimum and the index $[minf, indexMin] = \min(f)$.
- 3 Assign: $\gamma_{min} = r(indexMin)$.

Independent problem: Try writing it as a linear program!

The case $p = \infty$

Solution of minimize $\max_{0 \leq t \leq T_{max} - \tau} |Y(t + \tau) - \gamma V(t)|$ is given by the following linear γ program:

$$\begin{aligned} & \text{minimize} && s \\ & -s \leq Y(t + \tau) - \gamma V(t) \leq s, && 0 \leq t \leq T_{max} - \tau \end{aligned}$$

It can be rewritten into a standard form with vector $x = [s; \gamma]$, matrix A , vectors b , $f = [1; 0]$, lower bound $\mathbf{0} = [0; 0]$ and upper bound $u_\infty = [\infty; 1]$:

$$\begin{aligned} & \text{minimize} && f^T x \\ & Ax \leq b \\ & \mathbf{0} \leq x \leq u_\infty \end{aligned}$$

where:

$$A = \begin{bmatrix} -1 & -V(0) \\ -1 & V(0) \\ \vdots & \vdots \\ -1 & -V(T_{max} - \tau) \\ -1 & V(T_{max} - \tau) \end{bmatrix}, \quad b = \begin{bmatrix} -Y(\tau) \\ Y(\tau) \\ \vdots \\ -Y(T_{max}) \\ Y(T_{max}) \end{bmatrix}.$$

Note: A is a matrix of size $2(T_{max} - \tau + 1) \times 2$ and b is vector of length $2(T_{max} - \tau + 1)$.

SIR Model with Vitals

A simple modification of the SIR vanilla model is to consider vital signals, such as births and deaths at separate processes. In normalized form this becomes:

$$\begin{cases} \frac{ds}{dt} = \frac{\Lambda}{N} - \beta si - \mu s & , \quad s(0) = \frac{S_0}{N} \\ \frac{di}{dt} = \beta si - \alpha i - \mu i & , \quad i(0) = \frac{I_0}{N} \\ \frac{dr}{dt} = \alpha i - \mu r & , \quad r(0) = \frac{R_0}{N} \end{cases} \quad (\text{SIR Model})$$

where $\Lambda \geq 0$ is the constant source of births (=number of births/day) and $\mu \geq 0$ is the natural death rate (i.e., in the absence of this virus). Its reciprocal $1/\mu$ represents the average life expectancy.