STAT 770 Oct. 19 Lecture 14 Quasilikelihood and Overdispersion

Reading and Topics for this lecture: Secs. 4.4.1-4.4.5, 'quasi' family within glm in R, also Ch 13 through 13.1.2, Sec. 14.3-14.4

(1) Idea of Overdispersion – Important!

(2) Specific Models for Overdispersion, e.g. Negative Binomial

(3) Quasilikelihood estension of GLM

(4) R Code for Quasilikelihood fits

(5) Summarizing Predictive Power of Logistic Regression (Sec. 6.3)

Extra Heterogeneity among Observations

A key assumption required for validity of GLM methods is homogeneity and independence (*iid individual observations*).

What if observations in groups (i.e., **clusters**) share a random effect, i.e., observations are dependent among themselves ?

Examples:

(a) **Beta-Binomial** $Y_i \sim \text{Binom}(n_i, p_i), p_i \sim \text{Beta}(\alpha, \beta),$

(b) Negative-Binomial $Y_i \sim \text{Poisson}(\lambda_i), \ \lambda_i \sim \text{Gamma}(k, k/\mu)$

$$Var(Y_i) = E(Var(Y_i | \lambda_i)) + Var(E(Y_i | \lambda_i)) = E(\lambda_i) + Var(\lambda_i)$$

= $\mu + \mu^2/k$ over-dispersion

NegBin Model

(material covered in detail in Sec. 14.4) In example (b): Negative Binomial $(k, \mu/(\mu + k))$ distribution

$$P(Y_{i} = m) = \int_{0}^{\infty} \frac{(k/\mu)^{k}}{(k-1)!} x^{k-1} e^{-xk/\mu} e^{-x} \frac{x^{m}}{m!} dx = \frac{(k/\mu)^{k}}{\Gamma(k)\Gamma(m+1)} \int_{0}^{\infty} x^{k+m-1} e^{-x(1+k/\mu)} dx$$
$$= \frac{\Gamma(m+k)}{\Gamma(k)\Gamma(m+1)} (k/\mu)^{k} (1+k/\mu)^{-(m+k)}$$

Randomness in the Poisson parameter has the effect of changing the link! More generally, need estimation strategy to recognize overdispersion.

Quasilikelihood Extension of GLM

Assume instead of exponential family, $Y_i \sim h(y, \phi) e^{(\theta y - c(\theta))/a(\phi)}$ ϕ is a new dispersion parameter

Check again (same as before) that $c'(\theta) = E(Y_i)$. For variances:

$$0 = \frac{\partial}{\partial \theta} \int (y - c'(\theta)) f(y, \theta) \, dy = \int \left[\frac{(y - c'(\theta))^2}{a(\phi)} - c''(\theta) \right] f(y, \theta) \, dy$$

which implies: $\operatorname{Var}(Y) = a(\phi) c''(\theta)$

So score estimating equation (for θ) becomes

$$0 = \sum_{i=1}^{n} \nabla_{\beta} \left[(\theta_{i} Y_{i} - c(\theta_{i})) / a(\phi) \right] = \sum_{i=1}^{n} X_{i} \frac{Y_{i} - \mu_{i}}{g'(\mu_{i}) a(\phi) v(\mu_{i})}$$

same with factor $1/a(\phi)$ and $J = \mathbf{X}^{tr} (W/a(\phi)) \mathbf{X}$

Dispersion Parameter

 $a(\phi)$ generally found not from likelihood but from a moment expression like

$$n^{-1} \sum_{i=1}^{n} \frac{(Y_i - \hat{\mu}_i)^2}{v(\hat{\mu}_i)} = n^{-1} \sum_{i=1}^{n} \frac{(Y_i - g^{-1}(X_i^{tr} \hat{\beta})^2)}{v \circ g^{-1}(X_i^{tr} \hat{\beta})}$$

to indicate overdispersion when this estimate is clearly > 1

R Script with Illustrations of Methods

- (i) Quasilikelihood example (simulated), RscriptLec14.RLog(ii) Quasilikelihood example as in book with Crab mating, in script GLMsScript.RLog
- (iii) Predictive Power of Logistic Regression Models
 - -% reduction in sum of squared errors
 - -% reduction in deviance
 - ROC plot and AUC

Predictive Power of Logistic Regressions

Continue with this topic later on, using

- Binning of predicted values to make sure that the approximately correct proportions of Y = 1 cases are seen within each bin

- Cross-validated estimates of proportion of correct predictions

 Comparison with 'Machine Learning' algorithms for binary prediction