

DEPARTMENT OF MATHEMATICS  
UNIVERSITY OF MARYLAND  
GRADUATE WRITTEN EXAMINATION  
JANUARY, 2004

Statistics (Ph. D. Version)

*Instructions to the Student*

- a. Answer all six questions. Each will be graded from 0 to 10.
- b. Use a different booklet for each question. Write the problem number and your code number (**NOT YOUR NAME**) on the outside cover.
- c. Keep scratch work on separate pages in the same booklet.
- d. If you use a “well known” theorem in your solution to any problem, it is your responsibility to make clear which theorem you are using and to justify its use.

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**1.** Let  $\mathcal{P} = \{P_\theta\}$  be a family of distributions of a random element  $X \in \mathcal{X} \subset R^n$  parametrized by a discrete parameter  $\theta \in \{\theta_0, \theta_1, \theta_2\}$ . Assuming  $P_\theta$  is given by a density  $p(x; \theta)$  which is everywhere positive on  $\mathcal{X}$ , develop the minimal sufficient statistic for  $\theta$ .

**2.** Let  $(x_1, x_2, \dots, x_n)$  be a sample from a population  $X$  with density  $e^{-(x-\theta)}$ ,  $x \geq \theta$ . Find the uniformly minimum variance unbiased estimator of the value of the parameter function  $g(\theta) = P_\theta(1 < X < 2)$ .

**3.** Given a parameter  $\theta$ ,  $X_1, X_2, \dots, X_n$  are independent identically distributed random variables with  $E(X_i|\theta) = \theta$ ,  $\text{var}(X_i|\theta) = \sigma^2$ . The prior distribution of  $\theta$  has known mean  $a$  and variance  $b^2$ . The linear Bayesian estimator of  $\theta$ , by definition, minimizes the overall risk within the class of estimators of the form  $c_0 + \sum_{i=1}^n c_i X_i$  where the constant coefficients  $c_0, c_1, \dots, c_n$  do not depend upon the data. Find the linear Bayesian estimator of  $\theta$  with

respect to the squared-error loss function.

4. Let  $(x_1, x_2, \dots, x_n)$  be a sample from a population with density

$$f(x; \theta) = \frac{1}{B(3, 3)}(x - \theta)^2(1 - x + \theta)^2, \theta \leq x \leq 1 + \theta.$$

with  $\theta$  as a parameter. Here  $B(\alpha, \beta) = \int_0^1 x^{\alpha-1}(1-x)^{\beta-1}dx$  is the beta function,  $B(\alpha, \beta) = (\alpha-1)!(\beta-1)!/(\alpha+\beta-1)!$  for integer  $\alpha > 0, \beta > 0$ . Develop the method of moments estimator of  $\theta$  and calculate its efficiency.

5. Let  $(x_1, x_2, \dots, x_m)$  and  $(y_1, y_2, \dots, y_n)$  be two independent samples from normal populations with parameters  $(\mu_1, \sigma_1^2)$  and  $(\mu_2, \sigma_2^2)$ , respectively. Assuming  $\sigma_1^2 = 4\sigma^2, \sigma_2^2 = \sigma^2, \sigma^2$  unknown, develop a test of a constant level  $\alpha$  for all  $\sigma^2 > 0$  of the null hypothesis  $H_0 : \mu_1 = \mu_2$  versus the alternative  $H_1 : \mu_1 \neq \mu_2$ .

6. Let  $(x_1, y_1), \dots, (x_n, y_n)$  be a sample from a bivariate normal population with mean vector  $(\theta_1, \theta_2)$ , regarded as a parameter and covariance matrix  $I$ , the identity matrix. For the maximum likelihood estimator  $\hat{g}_n$  of the value of the parameter function  $g(\theta_1, \theta_2) = (\theta_1 - \theta_2)^2$ , find the nondegenerate limiting distribution of  $a_n(\hat{g}_n - g(\theta_1, \theta_2))$  as  $n \rightarrow \infty$  for a properly chosen  $a_n$ .