

DEPARTMENT OF MATHEMATICS  
UNIVERSITY OF MARYLAND  
GRADUATE WRITTEN EXAMINATION  
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Statistics (M. A. Version)

*Instructions to the Student*

- a. Answer all six questions. Each will be graded from 0 to 10.
- b. Use a different booklet for each question. Write the problem number and your code number (**NOT YOUR NAME**) on the outside cover
- c. Keep scratch work on separate pages in the same booklet.
- d. If you use a “well known” theorem in your solution to any problem, it is your responsibility to make clear which theorem you are using and to justify its use

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1. (i) Let  $p(x, y; \theta)$ ,  $\theta \in \Theta$  be a family of joint pdf's of random vectors  $X, Y$  with  $\theta$  as a parameter.

Assuming  $p(x, y; \theta) > 0$  for all  $x, y, \theta$ , show that if  $Y$  is sufficient for  $\theta$  and  $X$  and  $Y$  are independent when  $\theta = \theta_0 \in \Theta$ , then  $X$  and  $Y$  are independent for all  $\theta \in \Theta$ .

(ii) Let  $(X, Y)$  be a random vector having bivariate normal distribution with

$$E(X) = \theta_1, E(Y) = \theta_2, \text{var}(X) = \text{var}(Y) = 1, \text{corr}(X, Y) = \rho$$

with  $(\theta_1, \theta_2)$  as a parameter (and  $\rho$  known).

Using the factorization theorem, show that if the parameter set is  $\Theta = \{(\theta_1, \theta_2) : \theta_2 = \rho\theta_1\}$  then  $Y$  is sufficient for  $(\theta_1, \theta_2)$ .

2. Let  $(x_1, \dots, x_n)$  and  $(y_1, \dots, y_n)$  be independent samples from populations with pdf's  $f(x - \theta_1)$  and  $f(x - \theta_2)$ , respectively, with  $f(x) = e^{-x}$ ,  $x \geq 0$  and  $\theta_1, \theta_2$  as parameters.

(i) Find the MLE  $\hat{\Delta}_n$  of  $\Delta = \theta_1 - \theta_2$  and calculate  $E(\hat{\Delta}_n)$  and  $\text{var}(\hat{\Delta}_n)$ .

(ii) Show that the limiting distribution of  $n(\hat{\Delta}_n - \Delta)$  as  $n \rightarrow \infty$  is double exponential, i. e., given by the pdf  $(1/2)\lambda \exp\{-\lambda|x|\}$ ,  $x \in (-\infty, +\infty)$  for some  $\lambda > 0$ .

(Hint: Using the moment generating function may simplify the calculations.)

3. Let  $(x_1, \dots, x_n)$ ,  $(y_1, \dots, y_n)$ ,  $(z_1, \dots, z_n)$  be independent samples from exponential populations with densities

$$f(x; \lambda_1), f(y; \lambda_2), f(z; \lambda_3),$$

respectively, where  $f(u; \lambda) = (1/\lambda)e^{-u/\lambda}$ ,  $u > 0$  and  $\lambda_i > 0$ ,  $i = 1, 2, 3$  are parameters.

Construct the LR (Likelihood Ratio) test of level  $\alpha$  for testing

$$H_0 : \lambda_1 = \lambda_2 = \lambda_3 \text{ vs } H_1 : \lambda_i \neq \lambda_j \text{ for some } i, j$$

(ii) Assuming  $\lambda_2/\lambda_1 = c_1$ ,  $\lambda_3/\lambda_1 = c_2$ , show that the power function of the LR test depends only on  $c_1, c_2$ .

4. Given  $\theta$ ,  $(x_1, \dots, x_n)$  is a sample from a population with pdf  $e^{(x-\theta)}$ ,  $x \leq \theta$  (notice that the distribution is concentrated on  $(-\infty, \theta)$ ).

(i) Show that the family of pdf's

$$\pi(\theta; a, \lambda) = \lambda e^{-\lambda(\theta-a)}, \theta \geq a$$

with  $\lambda > 0$ ,  $a \in (-\infty, +\infty)$  as parameters is a conjugate family of prior pdf's.

(ii) Assuming the prior pdf  $\pi(\theta)$  belonging to the conjugate family, find the Bayes estimator of  $\theta$  for the quadratic loss function  $L(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^2$ .

5. Let  $(x_1, \dots, x_m)$  and  $(y_1, \dots, y_n)$  be independent samples of sizes  $m$  and  $n$  from populations with pdf's

$$f(x; \lambda_1) = (1/2)\lambda_1^3 x^2 e^{-\lambda_1 x}, x \geq 0$$

and

$$f(y; \lambda_2) = (1/6)\lambda_2^4 y^3 e^{-\lambda_2 y}, y \geq 0,$$

respectively, with  $\lambda_1, \lambda_2$  as parameters.

- (i) Based on the sufficient statistics, construct a pivot for  $\lambda_1/\lambda_2$ .
- (ii) Express the distribution of the pivot in terms of the  $F$ -distribution and construct a level  $1 - \alpha$  confidence interval for  $\lambda_1/\lambda_2$ .

6. Let  $(x_1, \dots, x_n)$  be a sample from a normal population  $N(\mu, \sigma^2)$  with parameters  $\mu, \sigma^2$ . The  $100(1 - \alpha)$ -th population percentile  $\eta_\alpha = \eta(\alpha)$  is a function of  $\mu, \sigma^2$ ,  $\eta_\alpha = \eta_\alpha(\mu, \sigma^2)$ .

- (i) Find the MLE  $\hat{\eta}_n(\alpha)$  of  $\eta(\alpha)$  and calculate its variance.
- (ii) Find the limiting distribution of  $\sqrt{n}(\hat{\eta}_n(\alpha) - \eta(\alpha))$  as  $n \rightarrow \infty$ .