

DEPARTMENT OF MATHEMATICS  
UNIVERSITY OF MARYLAND  
GRADUATE WRITTEN EXAMINATION  
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Statistics (Ph. D. Version)

*Instructions to the Student*

- a. Answer all six questions. Each will be graded from 0 to 10.
  - b. Use a different booklet for each question. Write the problem number and your code number (**NOT YOUR NAME**) on the outside cover
  - c. Keep scratch work on separate pages in the same booklet.
  - d. If you use a “well known” theorem in your solution to any problem, it is your responsibility to make clear which theorem you are using and to justify its use.
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1. Let  $\{p(x; \theta_1, \theta_2), (\theta_1, \theta_2) \in \Theta_1 \times \Theta_2\}$  be a family of pdf's on a measurable space  $(\mathcal{X}, \mathcal{A})$  with  $(\theta_1, \theta_2)$  as a parameter. Assume that there exists  $\theta_1^* \in \Theta_1$  such that  $T_1 = T_1(X)$  is sufficient for the family  $\{p(x; \theta_1^*, \theta_2), \theta_2 \in \Theta_2\}$ . Assume also that for any  $\tilde{\theta}_2 \in \Theta_2$ , a statistic  $T_2 = T_2(X)$  is sufficient for the family  $\{p(x; \theta_1, \tilde{\theta}_2), \theta_1 \in \Theta_1\}$ .

(i) Write the versions of the factorization theorem expressing the above properties.

(ii) Assuming  $p(x; \theta_1, \theta_2) > 0$  for all  $x, \theta_1, \theta_2$ , show that the statistic  $T = (T_1, T_2)$  is sufficient for  $(\theta_1, \theta_2)$  (i. e., for the family  $\{p(x; \theta_1, \theta_2), (\theta_1, \theta_2) \in \Theta_1 \times \Theta_2\}$ ).

2. Let  $(x_1, \dots, x_n)$  and  $(y_1, \dots, y_n)$  be independent samples from populations with pdf's  $f(x - \theta_1)$  and  $f(x - \theta_2)$ , respectively, with  $f(x) = e^{-x}$ ,  $x \geq 0$  and  $\theta_1, \theta_2$  as parameters.

(i) Find the MLE  $\hat{\Delta}_n$  of  $\Delta = \theta_1 - \theta_2$  and calculate  $E(\hat{\Delta}_n)$  and  $\text{var}(\hat{\Delta}_n)$ .

(ii) Find the normalizing constants  $a_n > 0$  and the explicit formula of the pdf of the nondegenerate limiting distribution of  $a_n(\hat{\Delta}_n - \Delta)$  as  $n \rightarrow \infty$ .

(**Hint:** Using the moment generating function may simplify the calculations.)

3. Let  $(x_1, \dots, x_n)$ ,  $(y_1, \dots, y_n)$ ,  $(z_1, \dots, z_n)$  be independent samples from exponential populations with densities

$$f(x; \lambda_1), f(y; \lambda_2), f(z; \lambda_3),$$

respectively, where  $f(u; \lambda) = (1/\lambda)e^{-u/\lambda}$ ,  $u > 0$  and  $\lambda_i > 0$ ,  $i = 1, 2, 3$  are parameters.

(i) Construct the LR (Likelihood Ratio) test of level  $\alpha$  for testing

$$H_0 : \lambda_1 = \lambda_2 = \lambda_3 \text{ vs } H_1 : \lambda_i \neq \lambda_j \text{ for some } i, j.$$

(ii) Show that the power function of the LR test depends only on  $\lambda_2/\lambda_1, \lambda_3/\lambda_1$ .

4. Given  $\theta$ ,  $(x_1, \dots, x_n)$  is a sample from a population with pdf  $e^{x-\theta}$ ,  $x \leq \theta$  (notice that the distribution is concentrated on  $(-\infty, \theta)$ )

(i) Construct the conjugate family of prior pdf's parameterized by a two-dimensional parameter.

(ii) Assuming the prior pdf  $\pi(\theta)$  belongs to the conjugate family, find the Bayes estimator of  $\theta$  for the quadratic loss function  $L(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^2$ .

5. Let  $(x_1, \dots, x_m)$  and  $(y_1, \dots, y_n)$  be independent samples of sizes  $m$  and  $n$  from populations with pdf's

$$f(x; \lambda_1) = (1/2)\lambda_1^3 x^2 e^{-\lambda_1 x}, \quad x \geq 0$$

and

$$f(y; \lambda_2) = (1/6)\lambda_2^4 y^3 e^{-\lambda_2 y}, \quad y \geq 0,$$

respectively, with  $\lambda_1, \lambda_2$  as parameters.

(i) Based on the sufficient statistics, construct a pivot for  $\lambda_1/\lambda_2$ .

(ii) Express the distribution of the pivot in terms of the  $F$ -distribution and construct a level  $1 - \alpha$  confidence interval for  $\lambda_1/\lambda_2$ .

6. Let  $(x_1, \dots, x_n)$  be a sample from a normal population  $N(\mu, \sigma^2)$  with parameters  $\mu, \sigma^2$ . The  $100(1 - \alpha)$ -th population percentile  $\eta_\alpha = \eta(\alpha)$  is a function of  $\mu, \sigma^2$ ,  $\eta_\alpha = \eta_\alpha(\mu, \sigma^2)$

(i) Find the MLE  $\hat{\eta}_n(\alpha)$  of  $\eta(\alpha)$  and show that it is biased

(ii) Find the limiting distribution of  $\sqrt{n}(\hat{\eta}_n(\alpha) - \eta(\alpha))$  as  $n \rightarrow \infty$